Long Horizon Stock Price Models or Stochastic Perpetuities

Dilip B. Madan

Robert H. Smith School of Business

Zurich Spring School 2015, April 3

• Remarks on long horizon modeling issues.

- Remarks on long horizon modeling issues.
- Modeling balanced markets.

- Remarks on long horizon modeling issues.
- Modeling balanced markets.
- The discounted variance gamma process.

- Remarks on long horizon modeling issues.
- Modeling balanced markets.
- The discounted variance gamma process.
- The structure of booms and busts.

- Remarks on long horizon modeling issues.
- Modeling balanced markets.
- The discounted variance gamma process.
- The structure of booms and busts.
- The discounted variance gamma with booms and busts.

- Remarks on long horizon modeling issues.
- Modeling balanced markets.
- The discounted variance gamma process.
- The structure of booms and busts.
- The discounted variance gamma with booms and busts.
- Calibration of discounted variance gamma and asymptotic implied volatilities.

- Remarks on long horizon modeling issues.
- Modeling balanced markets.
- The discounted variance gamma process.
- The structure of booms and busts.
- The discounted variance gamma with booms and busts.
- Calibration of discounted variance gamma and asymptotic implied volatilities.
- Valuation of path dependent stochastic perpetuities.

- Remarks on long horizon modeling issues.
- Modeling balanced markets.
- The discounted variance gamma process.
- The structure of booms and busts.
- The discounted variance gamma with booms and busts.
- Calibration of discounted variance gamma and asymptotic implied volatilities.
- Valuation of path dependent stochastic perpetuities.
- Valuation of coupon for stock above a barrier.

• Stock prices have been modeled, Williams (1938), as claims to future dividend distributions.

- Stock prices have been modeled, Williams (1938), as claims to future dividend distributions.
- Alternate approaches emphasize stock prices as claims to future free cash flows, Copeland, Koller and Murrin (1994).

- Stock prices have been modeled, Williams (1938), as claims to future dividend distributions.
- Alternate approaches emphasize stock prices as claims to future free cash flows, Copeland, Koller and Murrin (1994).
- Or future abnormal earnings Preinreich (1938), Edwards and Bell (1961) and Ohlson (1995).

- Stock prices have been modeled, Williams (1938), as claims to future dividend distributions.
- Alternate approaches emphasize stock prices as claims to future free cash flows, Copeland, Koller and Murrin (1994).
- Or future abnormal earnings Preinreich (1938), Edwards and Bell (1961) and Ohlson (1995).
- We infer from such observations that markets are comfortable in valuing claims with no clear distribution dates or explicit payoffs.

- Stock prices have been modeled, Williams (1938), as claims to future dividend distributions.
- Alternate approaches emphasize stock prices as claims to future free cash flows, Copeland, Koller and Murrin (1994).
- Or future abnormal earnings Preinreich (1938), Edwards and Bell (1961) and Ohlson (1995).
- We infer from such observations that markets are comfortable in valuing claims with no clear distribution dates or explicit payoffs.
- You get some random payoffs sometime and possibly well past your remaining lifetime.

Discounted stock prices in financial engineering models

• By way of contrast the derivative pricing literature, beginning with Black and Scholes (1973) and Merton (1973) models the discounted price of a non-dividend paying stock directly as a risk neutral martingale.

Discounted stock prices in financial engineering models

- By way of contrast the derivative pricing literature, beginning with Black and Scholes (1973) and Merton (1973) models the discounted price of a non-dividend paying stock directly as a risk neutral martingale.
- We ask here what kind of martingale should one be considering.

Discounted stock prices in financial engineering models

- By way of contrast the derivative pricing literature, beginning with Black and Scholes (1973) and Merton (1973) models the discounted price of a non-dividend paying stock directly as a risk neutral martingale.
- We ask here what kind of martingale should one be considering.
- In particular should it be uniformly integrable.

• Imagine all gains, be they dividends, free cash flows or abnormal earnings, being invested at market rates, to engineer no distributions, with the resulting accumulation being embedded into the final stock price at time *t*.

- Imagine all gains, be they dividends, free cash flows or abnormal earnings, being invested at market rates, to engineer no distributions, with the resulting accumulation being embedded into the final stock price at time *t*.
- In addition the stock at time t accounts for the value at time t of all gains in the future past time t.

- Imagine all gains, be they dividends, free cash flows or abnormal earnings, being invested at market rates, to engineer no distributions, with the resulting accumulation being embedded into the final stock price at time *t*.
- In addition the stock at time t accounts for the value at time t of all gains in the future past time t.
- Now discount back to time zero and let t go to infinity.

- Imagine all gains, be they dividends, free cash flows or abnormal earnings, being invested at market rates, to engineer no distributions, with the resulting accumulation being embedded into the final stock price at time *t*.
- In addition the stock at time t accounts for the value at time t of all gains in the future past time t.
- Now discount back to time zero and let t go to infinity.
- There will be no future left after time t but the present value computed will converge to the present value of all gains over all time as evaluated at infinity.

- Imagine all gains, be they dividends, free cash flows or abnormal earnings, being invested at market rates, to engineer no distributions, with the resulting accumulation being embedded into the final stock price at time *t*.
- In addition the stock at time t accounts for the value at time t of all gains in the future past time t.
- Now discount back to time zero and let t go to infinity.
- There will be no future left after time t but the present value computed will converge to the present value of all gains over all time as evaluated at infinity.
- The resulting discounted gains martingale converges at infinity to the present value of all gains through time and is thus a uniformly integrable martingale.

Discounted stock price in financial engineering models

• The discounted stock price in most financial engineering models is not a uniformly integrable martingale.

Discounted stock price in financial engineering models

- The discounted stock price in most financial engineering models is not a uniformly integrable martingale.
- In fact it converges to zero with probability one at infinity.

Discounted stock price in financial engineering models

- The discounted stock price in most financial engineering models is not a uniformly integrable martingale.
- In fact it converges to zero with probability one at infinity.
- We wish to consider uniformly integrable martingale models for the discounted stock price.

• We have a discounting process

$$Y(t) = \int_0^t e^{-r(u)} du$$

• We have a discounting process

$$Y(t) = \int_0^t e^{-r(u)} du$$

• And a positive martingale M(t) for the discounted stock price that we take to be the stochastic exponential of the martingale m^{Y} where

$$m^{Y}(t) = \int_{0}^{t} Y(u) dm(u)$$

• We have a discounting process

$$Y(t) = \int_0^t e^{-r(u)} du$$

• And a positive martingale M(t) for the discounted stock price that we take to be the stochastic exponential of the martingale m^{Y} where

$$m^{Y}(t) = \int_{0}^{t} Y(u) dm(u)$$

• The martingale m(t) represents the shocks to free cash flows or abnormal earnings.

• We have a discounting process

$$Y(t) = \int_0^t e^{-r(u)} du$$

• And a positive martingale M(t) for the discounted stock price that we take to be the stochastic exponential of the martingale m^{Y} where

$$m^{Y}(t) = \int_{0}^{t} Y(u) dm(u)$$

- The martingale m(t) represents the shocks to free cash flows or abnormal earnings.
- The discounted stock is the stochastic exponential of discounted shocks to earnings or free cash flows.

Explicit representation of positive martingale

• Suppose m(t) has a continuous martingale component $m^{(c)}(t)$ and finite variation jump compensator $\nu(dx, dt)$ with

$$m(t) = m^{(c)}(t) + \int_0^t \int_{-\infty}^\infty x \left(\mu(dx, du) - \nu(dx, du) \right)$$

Explicit representation of positive martingale

• Suppose m(t) has a continuous martingale component $m^{(c)}(t)$ and finite variation jump compensator $\nu(dx, dt)$ with

$$m(t) = m^{(c)}(t) + \int_0^t \int_{-\infty}^\infty x \left(\mu(dx, du) - \nu(dx, du) \right)$$

• Then

$$M(t) = \exp\left(\begin{array}{c}\int_0^t Y(u)dm(u) - \frac{1}{2}\int_0^t Y^2(u)d\langle m^c \rangle_u\\ -\int_0^t \int_{-\infty}^\infty \left(\exp\left(Y(u)x\right) - 1\right)\nu(dx, du)\end{array}\right).$$

Balanced Market Condition

• For a stock in a balanced market we ask that M be uniformly integrable or equivalently that $M(\infty)$ is a well defined terminal random variable with

 $M(t) = E_t [M(\infty)].$

Balanced Market Condition

• For a stock in a balanced market we ask that M be uniformly integrable or equivalently that $M(\infty)$ is a well defined terminal random variable with

$$M(t) = E_t [M(\infty)].$$

• The discounted stock price is then the conditional expectation of a stochastic claim defined at infinity.

Balanced Market Condition

• For a stock in a balanced market we ask that M be uniformly integrable or equivalently that $M(\infty)$ is a well defined terminal random variable with

$$M(t) = E_t [M(\infty)].$$

- The discounted stock price is then the conditional expectation of a stochastic claim defined at infinity.
- It is consistently valued through time in the market in line with its final resolution occuring at infinity.

Balanced Geometric Brownian Motion

• In the case of geometric Brownian motion with $m^c = \sigma W$ for a standard Brownian motion $W = (W(t), t \ge 0)$ and $Y(t) = e^{-rt}$ we have, $\langle m^c \rangle = \sigma^2 t$ and

$$M(t) = \exp\left(\int_0^t e^{-ru}\sigma dW(u) - \frac{\sigma^2}{2}\int_0^t e^{-2ru}du\right)$$

Balanced Geometric Brownian Motion

• In the case of geometric Brownian motion with $m^c = \sigma W$ for a standard Brownian motion $W = (W(t), t \ge 0)$ and $Y(t) = e^{-rt}$ we have, $\langle m^c \rangle = \sigma^2 t$ and

$$M(t) = \exp\left(\int_0^t e^{-ru} \sigma dW(u) - \frac{\sigma^2}{2} \int_0^t e^{-2ru} du\right)$$

• In this case $\log(M(\infty))$ is normal with mean $-\sigma^2/(4r)$ and variance $\sigma^2/(2r)$.
Balanced Geometric Brownian Motion

• In the case of geometric Brownian motion with $m^c = \sigma W$ for a standard Brownian motion $W = (W(t), t \ge 0)$ and $Y(t) = e^{-rt}$ we have, $\langle m^c \rangle = \sigma^2 t$ and

$$M(t) = \exp\left(\int_0^t e^{-ru} \sigma dW(u) - \frac{\sigma^2}{2} \int_0^t e^{-2ru} du\right)$$

- In this case $\log(M(\infty))$ is normal with mean $-\sigma^2/(4r)$ and variance $\sigma^2/(2r)$.
- For $t < \infty$, $\log(M(t))$ is normal with mean $-(\sigma^2/2)(1 \exp(-2rt))/(2r)$ and variance $\sigma^2(1 \exp(-2rt))/(2r)$.

Discounted Variance Gamma Model

• We take Y(u) to be $\exp(-ru)$ and m(u) to be the difference of two standard gamma processes

$$m(u) = \gamma_p(u) - \gamma_n(u)$$

Discounted Variance Gamma Model

• We take Y(u) to be $\exp(-ru)$ and m(u) to be the difference of two standard gamma processes

$$m(u) = \gamma_p(u) - \gamma_n(u)$$

• This yields

$$m_Y(t) = \int_0^t e^{-ru} d\gamma_p(u) - \int_0^t e^{-ru} d\gamma_n(u).$$

Discounted Variance Gamma (DVG) with Parameters

• Now Introduce parameters and allow the two gamma processes affecting positive and negative shocks to have independent mean and variance rates.

Discounted Variance Gamma (DVG) with Parameters

- Now Introduce parameters and allow the two gamma processes affecting positive and negative shocks to have independent mean and variance rates.
- This is accommodated by space scaling and a time change.

Discounted Variance Gamma (DVG) with Parameters

- Now Introduce parameters and allow the two gamma processes affecting positive and negative shocks to have independent mean and variance rates.
- This is accommodated by space scaling and a time change.
- Hence for parameters b_p , c_p , b_n , c_n we write for $X = (X(t), t \ge 0)$ the process

$$X(t) = \int_0^t b_p e^{-ru} d\gamma_p(c_p u) - \int_0^t b_n e^{-ru} d\gamma_n(c_n u)$$

=
$$\int_0^{c_p t} b_p e^{-ru/c_p} d\gamma_p(u) - \int_0^{c_n t} b_n e^{-ru/c_n} d\gamma_n(u)$$

=
$$H_p(t) - H_n(t)$$

Compensator for DVG

• The process m(t) is given by the jump compensator

$$\nu(dx, dt) = \left(\mathbf{1}_{x<0}c_n \frac{\exp(-|x|/b_n)}{|x|} + \mathbf{1}_{x>0}c_p \frac{\exp(-x/b_p)}{x}\right) dx dt.$$

Characteristic function for DVG

• The characteristic function for X(t) is given by

$$\phi_{X(t)}(u) = \exp\left(\begin{array}{c} \frac{c_p}{r} \left(dilog \left(iub_p\right) - dilog \left(iub_p e^{-rt}\right)\right) \\ + \frac{c_n}{r} \left(dilog \left(-iub_n\right) - dilog \left(-iub_n e^{-rt}\right)\right) \end{array}\right)$$

Zurich Spring School 2015, April 3

14

Characteristic function for DVG

• The characteristic function for X(t) is given by

$$\phi_{X(t)}(u) = \exp\left(\begin{array}{c} \frac{c_p}{r} \left(dilog \left(iub_p\right) - dilog \left(iub_p e^{-rt}\right)\right) \\ + \frac{c_n}{r} \left(dilog \left(-iub_n\right) - dilog \left(-iub_n e^{-rt}\right)\right) \end{array}\right)$$

• The *dilog* function is given by

$$dilog(x) = -\int_0^x \frac{\ln(1-t)}{t} dt.$$

• The limiting random variable $X(\infty)$ has characteristic function

$$\phi_{X(\infty)}(u) = \exp\left(\frac{c_p}{r}dilog\left(iub_p\right) + \frac{c_n}{r}dilog\left(-iub_n\right)\right).$$

• The limiting random variable $X(\infty)$ has characteristic function

$$\phi_{X(\infty)}(u) = \exp\left(\frac{c_p}{r}dilog\left(iub_p\right) + \frac{c_n}{r}dilog\left(-iub_n\right)\right).$$

• Using the analysis of Barndorff-Nielsen and Shepard (2002), one may relate this limiting random variable to the stationary distribution for the process of abnormal gains G(t) where we define

dG = -rGdt + dL

• The limiting random variable $X(\infty)$ has characteristic function

$$\phi_{X(\infty)}(u) = \exp\left(\frac{c_p}{r}dilog\left(iub_p\right) + \frac{c_n}{r}dilog\left(-iub_n\right)\right).$$

 Using the analysis of Barndorff-Nielsen and Shepard (2002), one may relate this limiting random variable to the stationary distribution for the process of abnormal gains G(t) where we define

$$dG = -rGdt + dL$$

• The driving process L(t) for undiscounted shocks is here given by

$$L(t) = b_p \gamma_p(c_p t) - b_n \gamma_n(c_n t).$$

Dilip B. Madan (Robert H. Smith School of

• The limiting random variable $X(\infty)$ has characteristic function

$$\phi_{X(\infty)}(u) = \exp\left(rac{c_p}{r}dilog\left(iub_p
ight) + rac{c_n}{r}dilog\left(-iub_n
ight)
ight).$$

• Using the analysis of Barndorff-Nielsen and Shepard (2002), one may relate this limiting random variable to the stationary distribution for the process of abnormal gains G(t) where we define

$$dG = -rGdt + dL$$

• The driving process L(t) for undiscounted shocks is here given by

$$L(t) = b_p \gamma_p(c_p t) - b_n \gamma_n(c_n t).$$

• The distribution of the random variable $X(\infty)$ is the stationary distribution for G(t) .

Exponential DVG

• For the exponential convexity correction define

$$\omega(t) = -\ln\left(\phi_{X(t)}(-i)\right).$$

Zurich Spring School 2015, April 3

16 50

Exponential DVG

• For the exponential convexity correction define

$$\omega(t) = -\ln\left(\phi_{X(t)}(-i)\right).$$

• Equivalently

$$\exp(\omega(t)) = rac{1}{E\left[\exp(X(t))
ight]}.$$

Exponential DVG

• For the exponential convexity correction define

$$\omega(t) = -\ln\left(\phi_{X(t)}(-i)\right).$$

Equivalently

$$\exp\left(\omega(t)
ight) = rac{1}{E\left[\exp(X(t))
ight]}.$$

• Given that X(t) is a process of independent but inhomogeneous increments we get that $M = (M(t), t \ge 0)$ is a martingale where

$$M(t) = \exp \left(X(t) + \omega(t)
ight)$$
 .

Zurich Spring School 2015, April 3

Dilip B. Madan (Robert H. Smith School of

• We model our risk neutral discounted stock price V as the four parameter martingale M.

- We model our risk neutral discounted stock price V as the four parameter martingale M.
- This model may be calibrated to option price data using standard Fourier transform methods of Carr and Madan (1999) from the characteristic function for X(t).

- We model our risk neutral discounted stock price V as the four parameter martingale M.
- This model may be calibrated to option price data using standard Fourier transform methods of Carr and Madan (1999) from the characteristic function for X(t).
- More specifically the stock price at maturity t is modeled as

 $S(t) = S_0 \exp\left((\widetilde{r}(t) - q(t))t + \omega(t) + X(t)\right)$

- We model our risk neutral discounted stock price V as the four parameter martingale M.
- This model may be calibrated to option price data using standard Fourier transform methods of Carr and Madan (1999) from the characteristic function for X(t).
- More specifically the stock price at maturity t is modeled as

$$S(t) = S_0 \exp\left((\widetilde{r}(t) - q(t))t + \omega(t) + X(t)\right)$$

• $\tilde{r}(t)$ and q(t) are the annualized discount rates and dividend yields for maturity t consistent with the forward price for maturity t.

Unbalanced Markets

• For generality consider separate discounting functions for abnormal gains and losses whereby we write for the positive and negative components of the log price the processes

$$X_p(t) = \int_0^t h_p(u) d\gamma^p(u)$$

$$X_n(t) = \int_0^t h_n(u) d\gamma^n(u).$$

Unbalanced Markets

• For generality consider separate discounting functions for abnormal gains and losses whereby we write for the positive and negative components of the log price the processes

$$X_p(t) = \int_0^t h_p(u) d\gamma^p(u)$$

$$X_n(t) = \int_0^t h_n(u) d\gamma^n(u).$$

• We may lose uniformly integrability by the stochastic exponential of compensated jumps failing to be uniformly integrable on either the positive or negative side or both.

Unbalanced Markets

• For generality consider separate discounting functions for abnormal gains and losses whereby we write for the positive and negative components of the log price the processes

$$X_p(t) = \int_0^t h_p(u) d\gamma^p(u)$$

$$X_n(t) = \int_0^t h_n(u) d\gamma^n(u).$$

- We may lose uniformly integrability by the stochastic exponential of compensated jumps failing to be uniformly integrable on either the positive or negative side or both.
- For the positive side we have a boom market and for the negative side we have a crashing market or one in a bust.

Conditions for uniform integrability

• In the context of Wiener gamma integrals the required integrability conditions may be shown to be

$$\int_0^\infty \log\left(1+h_i(u)\right) du < \infty, \quad i=p, n.$$

$$-\int_0^\infty \log\left(1-h_p(u)\right) du < \infty.$$

Conditions for uniform integrability

• In the context of Wiener gamma integrals the required integrability conditions may be shown to be

$$\int_0^\infty \log\left(1+h_i(u)\right) du < \infty, \quad i=p, n$$
$$-\int_0^\infty \log\left(1-h_p(u)\right) du < \infty.$$

• The last condition is needed to ensure a finite exponential moment for X_p .

The structure of booms

• Suppose the positive shocks fail to be uniformly integrable when exponentially compensated.

The structure of booms

- Suppose the positive shocks fail to be uniformly integrable when exponentially compensated.
- The integrability condition then fails for h_p and

$$\int_0^\infty \log\left(1+h_p(u)\right) du = \infty.$$

The structure of booms

- Suppose the positive shocks fail to be uniformly integrable when exponentially compensated.
- The integrability condition then fails for h_p and

J

$$\int_0^\infty \log\left(1+h_p(u)\right) du = \infty.$$

• It follows also that

$$\int_0^\infty h_p(u)du = \infty$$

Drift action for booms

• The process $L_p(t)$ for the logarithm of the positive shocks may then be defined as

$$L_p(t) = \int_0^t h_p(u) d\gamma^p(u) + \int_0^t \log(1-h_p(u)) du.$$

Drift action for booms

 The process L_p(t) for the logarithm of the positive shocks may then be defined as

$$L_p(t) = \int_0^t h_p(u) d\gamma^p(u) + \int_0^t \log(1 - h_p(u)) du.$$

• The expectation of $L_p(t)$ is

$$E[L_p(t)] = \int_0^t h_p(u) du + \int_0^t \log(1 - h_p(u)) du,$$

and as

$$-\log(1-x) > x$$

the expectation of $L_p(t)$ is negative and tending to negative infinity.

Summary for booms

Summary for booms

• The drift dominates and the logarithm tends to negative infinity.

Dilip B. Madan (Robert H. Smith School of

Zurich Spring School 2015, April 3

22 /

Summary for booms

Summary for booms

- The drift dominates and the logarithm tends to negative infinity.
- The terminal random variable is then zero.

The structure of busts

• The integrability condition now fails for h_n and we have

$$\int_0^\infty \log\left(1+h_n(u)\right) du = \infty$$

Zurich Spring School 2015, April 3

23

The structure of busts

• The integrability condition now fails for h_n and we have

$$\int_0^\infty \log\left(1+h_n(u)\right) du = \infty.$$

• It follows that

$$\int_0^\infty h_n(u)du=\infty,$$

The structure of busts

• The integrability condition now fails for h_n and we have

$$\int_0^\infty \log\left(1+h_n(u)\right) du = \infty.$$

It follows that

$$\int_0^\infty h_n(u)du=\infty,$$

• The process $L_n(t)$ for the logarithm of the negative shocks may be defined as

$$L_n(t) = -\int_0^t h_n(u)d\gamma^n(u) + \int_0^t \log(1+h_n(u))du.$$

Jump action for busts

• The expectation of the logarithm of the negative shocks is

$$E[L_n(t)] = -\int_0^t h_n(u) du + \int_0^t \log(1 + h_n(u)) du$$

Zurich Spring School 2015, April 3

24
Jump action for busts

• The expectation of the logarithm of the negative shocks is

$$E[L_n(t)] = -\int_0^t h_n(u) du + \int_0^t \log(1 + h_n(u)) du$$

• The inequality

 $\log(1+x) < x$

Jump action for busts

• The expectation of the logarithm of the negative shocks is

$$E[L_n(t)] = -\int_0^t h_n(u) du + \int_0^t \log(1 + h_n(u)) du$$

The inequality

$$\log(1+x) < x$$

• yields that this expectation is negative and tending to negative infinity.

Jump action for busts

• The expectation of the logarithm of the negative shocks is

$$E[L_n(t)] = -\int_0^t h_n(u) du + \int_0^t \log(1 + h_n(u)) du$$

The inequality

$$\log(1+x) < x$$

- yields that this expectation is negative and tending to negative infinity.
- However, this time the negative shocks are too large with insufficient compensation to keep the price from dropping.

Differences between booms and busts

• With regard to *booms* and *busts we observe* that a jump of size x > 0 is compensated in the exponential by a drift of $-(e^x - 1)$ and this dominates the jump with $(e^x - 1) > x$.

Differences between booms and busts

- With regard to *booms* and *busts we observe* that a jump of size x > 0 is compensated in the exponential by a drift of $-(e^x 1)$ and this dominates the jump with $(e^x 1) > x$.
- Similarly for a negative jump of size −x < 0 the compensation in the drift is (1 − e^{-x}) < x and the jumps dominate the drift compensation.

Differences between booms and busts

- With regard to *booms* and *busts we observe* that a jump of size x > 0 is compensated in the exponential by a drift of $-(e^x 1)$ and this dominates the jump with $(e^x 1) > x$.
- Similarly for a negative jump of size -x < 0 the compensation in the drift is $(1 e^{-x}) < x$ and the jumps dominate the drift compensation.
- Hence risk neutrally *booms* are ended by the negative drift while jumps dominate the positive drift in *busts*.

DVGBB generalization for booms and busts

• The *DVG* model is extended to accomodate booms and busts to form the model *DVGBB*.

DVGBB generalization for booms and busts

- The *DVG* model is extended to accomodate booms and busts to form the model *DVGBB*.
- With a view to accomodating *booms* and *busts* we consider the Wiener-Gamma integral

$$K(t) = \int_0^t (a + be^{-ru}) d\gamma(cu)$$
$$= \int_0^{ct} (a + be^{-ru/c}) d\gamma(u)$$

Dilip B. Madan (Robert H. Smith School of

Laplace transform of K(t)

• The Laplace transform of K(t) is given by

$$E\left[\exp\left(-\lambda K(t)\right)\right] = \exp\left(\frac{c}{r}\left(\Xi\left(\frac{1}{a+b}, a, \lambda\right) - \Xi\left(\frac{1}{a+be^{-rt}}, a, \lambda\right)\right)$$

where

$$\begin{split} \Xi\left(z,a,\lambda\right) &= \int \frac{\log\left(1+\frac{\lambda}{z}\right)}{z-az^2} dz \\ &= -dilog\left(\frac{a(\lambda+z)}{a\lambda+1}\right) \\ &-\log\left(\lambda+z\right)\log\left(1-\frac{a(\lambda+z)}{a\lambda+1}\right) + dilog(az) \\ &+\log(z)\log(1-az) - dilog\left(-\frac{z}{\lambda}\right) \\ &+\log(z)\log(\lambda) \\ &-\frac{1}{2}\log^2(z). \end{split}$$

Dilip B. Madan (Robert H. Smith School of

DVG calibrated to S&P 500 index options

• The discount rate for the response of the discounted stock price to abnormal earnings as captured in the processes H_p , H_n is taken to be fixed at a long term interest rate and in the reported calibrations this was set at the 30 year discount rate.

DVG calibrated to S&P 500 index options

- The discount rate for the response of the discounted stock price to abnormal earnings as captured in the processes H_p , H_n is taken to be fixed at a long term interest rate and in the reported calibrations this was set at the 30 year discount rate.
- We first fit the DVG model to 582 options covering 30 maturities for S&P 500 options as at market close on April 18 2012.

Calibration Result

• The result is as follows, where the fit statistics *rmse*, *aae* and *ape* are respectively, the root mean square error, the average absolute error and the average percentage error defined as the *aae* relative to the average option price in the data set.

r	=	0.02966
b _p	=	0.0066
c _p	=	512.9406

$$b_n = 122.4987$$

$$c_n = 0.0095$$

ape =
$$0.0311$$



• We observe that excluding some deep out-of-the-money short maturity options the fit is quite good for a four parameter model fit across such a large number of maturities.

- We observe that excluding some deep out-of-the-money short maturity options the fit is quite good for a four parameter model fit across such a large number of maturities.
- In general an exponential Lévy model does not fit options across maturity as skewness and excess kurtosis fall theoretically too sharply for such models.

- We observe that excluding some deep out-of-the-money short maturity options the fit is quite good for a four parameter model fit across such a large number of maturities.
- In general an exponential Lévy model does not fit options across maturity as skewness and excess kurtosis fall theoretically too sharply for such models.
- The performance of the variance gamma on this same data set is reported on later in the paper.

- We observe that excluding some deep out-of-the-money short maturity options the fit is quite good for a four parameter model fit across such a large number of maturities.
- In general an exponential Lévy model does not fit options across maturity as skewness and excess kurtosis fall theoretically too sharply for such models.
- The performance of the variance gamma on this same data set is reported on later in the paper.
- The four parameter Sato process (Carr, Geman, Madan and Yor (2007)) does fit the option surface but has a time inhomogeneity that is financially not as well motivated as the inhomogeneity for the DVG model.

- We observe that excluding some deep out-of-the-money short maturity options the fit is quite good for a four parameter model fit across such a large number of maturities.
- In general an exponential Lévy model does not fit options across maturity as skewness and excess kurtosis fall theoretically too sharply for such models.
- The performance of the variance gamma on this same data set is reported on later in the paper.
- The four parameter Sato process (Carr, Geman, Madan and Yor (2007)) does fit the option surface but has a time inhomogeneity that is financially not as well motivated as the inhomogeneity for the DVG model.
- Furthermore the Sato process generally cannot be extended to the larger maturities.

- We observe that excluding some deep out-of-the-money short maturity options the fit is quite good for a four parameter model fit across such a large number of maturities.
- In general an exponential Lévy model does not fit options across maturity as skewness and excess kurtosis fall theoretically too sharply for such models.
- The performance of the variance gamma on this same data set is reported on later in the paper.
- The four parameter Sato process (Carr, Geman, Madan and Yor (2007)) does fit the option surface but has a time inhomogeneity that is financially not as well motivated as the inhomogeneity for the DVG model.
- Furthermore the Sato process generally cannot be extended to the larger maturities.
- Eventually the time scaling involved will lead to a loss of exponential moments for the scaled unit time random variable Zurich Spring School 2015, April 3

Dilip B. Madan (Robert H. Smith School of I

Stochastic Perpetuities

Asymptotic Implied Volatilities

Asymptotic implied volatilities

• In addition we now have a well defined terminal random variable and we present the asymptotic implied volatility curve for options written on this terminal random variable in a zero rate economy.

Dilip B. Madan (Robert H. Smith School of I

Asymptotic Implied Volatilities



31 / 50

DVG Asymptotic Implied Vols SPX 20120418 30 maturities

r

Calibration of DVGBB

• The *DVGBB* model was calibrated to the 582 options across 30 maturities for the S&P 500 index as at April 18 2012. The results were as follows.

a_p	=	.0061415
b_p	=	.0003486
с _р	=	512.9436
a _n	=	.0112195
b _n	=	122.4965
c _n	=	.0089424
mse	=	4.3626
aae	=	3.6366
ape	_	.02781

DVGBB Calibration Remarks

• The calibration with boom and bust introduces the presence of some exposure to booms and busts making a 12% improvement in the root mean square error.

DVGBB Calibration Remarks

- The calibration with boom and bust introduces the presence of some exposure to booms and busts making a 12% improvement in the root mean square error.
- In general statistical tests are not available for calibration exercises.

DVGBB Calibration Remarks

- The calibration with boom and bust introduces the presence of some exposure to booms and busts making a 12% improvement in the root mean square error.
- In general statistical tests are not available for calibration exercises.
- Recently Madan (2013) has developed such procedures and future research could implement such significance tests for the boom and bust coefficients a_p , a_n .

Valuing path dependent stochastic perpetuities

• We consider claims that depend on the entire path of our martingale $(M(t), t \ge 0)$.

Valuing path dependent stochastic perpetuities

- We consider claims that depend on the entire path of our martingale $(M(t), t \ge 0)$.
- Consider by way of example a claim to the terminal cash flow

$$C(\infty) = (M(\infty) - K)^{+} + N \int_{0}^{\infty} e^{-rt} \mathbf{1}_{M(t) > a} dt$$

which is a call option on the terminal value of the discounted stock with strike K coupled with a note with notional N that is a perpetual coupon earned while the discounted stock is above a lower barrier.

Valuing path dependent stochastic perpetuities

- We consider claims that depend on the entire path of our martingale $(M(t), t \ge 0)$.
- Consider by way of example a claim to the terminal cash flow

$$C(\infty) = (M(\infty) - K)^+ + N \int_0^\infty e^{-rt} \mathbf{1}_{M(t)>a} dt$$

which is a call option on the terminal value of the discounted stock with strike K coupled with a note with notional N that is a perpetual coupon earned while the discounted stock is above a lower barrier.

• Such products are now popularly issued by numerous entities in the financial sector and are attractive to the investment community seeking principal protected yield enhancement opportunities.

Time zero and time t values

• In valuing this claim through time we are interested in the value in time zero dollars of

 $w(t)=E_t\left[C(\infty)\right].$

Time zero and time t values

• In valuing this claim through time we are interested in the value in time zero dollars of

$$w(t) = E_t [C(\infty)].$$

• The value in time t dollars is then given by

 $V(t)=e^{rt}w(t).$

Conditioning on time t

• To evaluate this conditional expectation we isolate the dependence of the claim on information known at time *t*, in particular the value X(t).

Conditioning on time t

- To evaluate this conditional expectation we isolate the dependence of the claim on information known at time *t*, in particular the value X(t).
- Now we know that

$$M(\infty) = \exp (X(\infty) + \omega(\infty)),$$

$$X(\infty) = \int_0^\infty b_p e^{-ru} d\gamma_p(c_p u) - \int_0^\infty b_n e^{-ru} d\gamma_n(c_n u)$$

Conditioning on time t

- To evaluate this conditional expectation we isolate the dependence of the claim on information known at time *t*, in particular the value X(t).
- Now we know that

$$M(\infty) = \exp(X(\infty) + \omega(\infty)),$$

$$X(\infty) = \int_0^\infty b_p e^{-ru} d\gamma_p(c_p u) - \int_0^\infty b_n e^{-ru} d\gamma_n(c_n u)$$

• By construction we have that

$$X(\infty) = X(t) + \int_{t}^{\infty} b_{p} e^{-ru} d\gamma_{p}(c_{p}u) - \int_{t}^{\infty} b_{n} e^{-ru} d\gamma_{n}(c_{n}u)$$

= $X(t) + e^{-rt} Y$

where the law of Y is that of $X(\infty)$.

The law of Y in the interim

• Furthermore we also have for u > t

$$X(u) = X(t) + e^{-rt}Y(u-t)$$

Zurich Spring School 2015, April 3

37 /
The law of Y in the interim

• Furthermore we also have for u > t

$$X(u) = X(t) + e^{-rt}Y(u-t)$$

• The law of Y(u) is that of

$$\int_0^u b_p e^{-rs} d\gamma_p(c_p s) - \int_0^u b_n e^{-rs} d\gamma_n(c_n s)$$

37

Evaluating the conditional expectation

• We then have

$$E_{t} [C(\infty)]$$

$$= E_{t} \begin{bmatrix} (\exp(X(\infty) + \omega(\infty)) - K)^{+} \\ +N \int_{0}^{\infty} e^{-rs} \mathbf{1}_{X(s)+\omega(s)>\ln(a)} ds \end{bmatrix}$$

$$= E_{t} \begin{bmatrix} (\exp(X(t) + e^{-rt}Y + \omega(\infty)) - K)^{+} + \\ \begin{pmatrix} N \int_{0}^{t} e^{-rs} \mathbf{1}_{X(s)>\ln(a)-\omega(s)} ds \\ +N \int_{t}^{\infty} e^{-rs} \mathbf{1}_{X(t)+e^{-rt}Y(s-t)+\omega(s)>\ln(a)} ds \end{pmatrix} \end{bmatrix}$$

$$= E_{t} \begin{bmatrix} (\exp(X(t) + e^{-rt}Y + \omega(\infty)) - K)^{+} + \\ \begin{pmatrix} N \int_{0}^{t} e^{-rs} \mathbf{1}_{X(s)>\ln(a)-\omega(s)} ds \\ +e^{-rt}N \int_{0}^{\infty} e^{-ru} \mathbf{1}_{Y(u)>(\ln(a)-\omega(t+u)-X(t))e^{rt}} du \end{pmatrix} \end{bmatrix}$$

38 50

Introduce the path statistic

• Define the path statistic

$$c(t) = \int_0^t e^{-rs} \mathbf{1}_{X(s) > \ln(a) - \omega(s)} ds.$$

Zurich Spring School 2015, April 3

39 50

Introduce the path statistic

• Define the path statistic

$$c(t)=\int_0^t e^{-rs}\mathbf{1}_{X(s)>\ln(a)-\omega(s)}ds.$$

• Given X(t), e^{-rt} , c(t) we may employ the terminal density of Y, f(y) and the densities g(y, u) for the random variables Y(u) to construct the above expectation as a function

$$w(t) = H(X, e^{-rt}) + Nc(t)$$

• We also know that w(t) is a martingale and furthermore $H(X(\infty), 0) + Nc(\infty) = (\exp(X(\infty) + \omega(\infty)) - K)^{+} + Nc(\infty).$

- We also know that w(t) is a martingale and furthermore $H(X(\infty), 0) + Nc(\infty) = (\exp(X(\infty) + \omega(\infty)) - K)^{+} + Nc(\infty).$
- We may then deduce the boundary condition

 $H(X,0) = (\exp(X + \omega(\infty)) - K)^+$

- We also know that w(t) is a martingale and furthermore $H(X(\infty), 0) + Nc(\infty) = (\exp(X(\infty) + \omega(\infty)) - K)^{+} + Nc(\infty).$
- We may then deduce the boundary condition

$$H(X,0) = (\exp(X + \omega(\infty)) - K)^+$$

• The martingale condition yields

$$-re^{-rt}H_{u} + e^{-rt}\mathbf{1}_{X > \ln(a) - \omega(t)}$$

+
$$\int_{-\infty}^{\infty} (H(X + y, e^{-rt}) - H(X, e^{-rt}))k(y)dy$$

= 0

- We also know that w(t) is a martingale and furthermore $H(X(\infty), 0) + Nc(\infty) = (\exp(X(\infty) + \omega(\infty)) - K)^{+} + Nc(\infty).$
- We may then deduce the boundary condition

$$H(X,0) = (\exp{(X + \omega(\infty))} - K)^+$$

• The martingale condition yields

$$-re^{-rt}H_{u} + e^{-rt}\mathbf{1}_{X>\ln(a)-\omega(t)}$$

+
$$\int_{-\infty}^{\infty} (H(X+y, e^{-rt}) - H(X, e^{-rt}))k(y)dy$$

= 0

• or in terms of *u*

$$-ruH_{u}(X, u) + u\mathbf{1}_{X > \ln(a) - \omega(-\frac{1}{r}\ln(u))}$$
$$+ \int_{-\infty}^{\infty} (H(X + y, u) - H(X, u))k(y)dy$$

April

Dilip B. Madan (Robert H. Smith School of

Stochastic Perpetuities

Direct computation of H(X,u)

• We may also directly compute the value as

$$H(X, u)$$

$$= E_{Y} \left[\left(\exp \left(X + uY + \omega(\infty) \right) - K \right)^{+} \right]$$

$$+ E_{(Y(s),s \ge 0)} \left[uN \int_{0}^{\infty} e^{-rs} \mathbf{1}_{Y(s) > (\ln(a) - \omega(-\frac{\ln(u)}{r} + s) - X)/u} ds \right]$$

Direct computation of H(X,u)

• We may also directly compute the value as

$$H(X, u)$$

$$= E_{Y} \left[\left(\exp \left(X + uY + \omega(\infty) \right) - K \right)^{+} \right]$$

$$+ E_{(Y(s),s \ge 0)} \left[uN \int_{0}^{\infty} e^{-rs} \mathbf{1}_{Y(s) > (\ln(a) - \omega(-\frac{\ln(u)}{r} + s) - X)/u} ds \right]$$

• We write in terms of the law of $Y(s) + \omega(s)$ or $Y + \omega(\infty)$ as follows $H(X, u) = E\left[(\exp\left(X + u(Y + \omega(\infty)) + (1 - u)\omega(\infty)\right) - K)\right]^{+} + E_{(Y(s) + \omega(s), s \ge 0)} \left[uN \int_{0}^{\infty} e^{-rs} \mathbf{1}_{Y(s) + \omega(s) > (\ln(a) - \omega(-\frac{\ln(u)}{r} + s) + u\omega(s))}\right]$

Direct computation of H(X,u)

We may also directly compute the value as

$$H(X, u)$$

$$= E_{Y} \left[\left(\exp \left(X + uY + \omega(\infty) \right) - K \right)^{+} \right]$$

$$+ E_{(Y(s), s \ge 0)} \left[uN \int_{0}^{\infty} e^{-rs} \mathbf{1}_{Y(s) > (\ln(a) - \omega(-\frac{\ln(u)}{r} + s) - X)/u} ds \right]$$

• We write in terms of the law of $Y(s) + \omega(s)$ or $Y + \omega(\infty)$ as follows H(X, u) $= E \left[(\exp \left(X + u(Y + \omega(\infty)) + (1 - u)\omega(\infty) \right) - K) \right]^+$ $+E_{(Y(s)+\omega(s),s\geq 0)}\left[uN\int_{0}^{\infty}e^{-rs}\mathbf{1}_{Y(s)+\omega(s)>(\ln(a)-\omega(-\frac{\ln(u)}{r}+s)+u\omega(s))}\right]$

 In this way we may compute the values of claims through time that are defined at infinity and in the interim have no cash flows but just a market value process. Zurich Spring School 2015, April 3

Dilip B. Madan (Robert H. Smith School of)

• There are now a number of yield enhancement securities offering coupon for the time spent by a stock index above a barrier.

- There are now a number of yield enhancement securities offering coupon for the time spent by a stock index above a barrier.
- For example Morgan Stanley issued such a note June 27, 2011, maturing June 30, 2031 and paying a 7.5% coupon provided the S&P 500 index was above 900 (See Registration Statement No. 333-156423, Rule 424(b) (2), Pricing Supplement No. 837).

- There are now a number of yield enhancement securities offering coupon for the time spent by a stock index above a barrier.
- For example Morgan Stanley issued such a note June 27, 2011, maturing June 30, 2031 and paying a 7.5% coupon provided the S&P 500 index was above 900 (See Registration Statement No. 333-156423, Rule 424(b) (2), Pricing Supplement No. 837).
- The notes can be quite long dated with maturities up to 20 or 30 years, and we take them here to be perpetuities.

- There are now a number of yield enhancement securities offering coupon for the time spent by a stock index above a barrier.
- For example Morgan Stanley issued such a note June 27, 2011, maturing June 30, 2031 and paying a 7.5% coupon provided the S&P 500 index was above 900 (See Registration Statement No. 333-156423, Rule 424(b) (2), Pricing Supplement No. 837).
- The notes can be quite long dated with maturities up to 20 or 30 years, and we take them here to be perpetuities.
- The payoff in present value terms is then

$$c\int_0^\infty e^{-rt}\mathbf{1}_{S(t)>a}dt.$$

- There are now a number of yield enhancement securities offering coupon for the time spent by a stock index above a barrier.
- For example Morgan Stanley issued such a note June 27, 2011, maturing June 30, 2031 and paying a 7.5% coupon provided the S&P 500 index was above 900 (See Registration Statement No. 333-156423, Rule 424(b) (2), Pricing Supplement No. 837).
- The notes can be quite long dated with maturities up to 20 or 30 years, and we take them here to be perpetuities.
- The payoff in present value terms is then

$$c\int_0^\infty e^{-rt}\mathbf{1}_{S(t)>a}dt.$$

• Alternatively in terms of the discounted stock martingale we could write

$$c\int_0^\infty e^{-rt}\mathbf{1}_{M(t)>ae^{-rt}}dt.$$

A GBM computation

• Suppose we take for our martingale the classical geometric Brownian motion model whereby

$$M_{C}(t) = e^{\sigma W(t) - \frac{\sigma^{2}t}{2}}$$

A GBM computation

• Suppose we take for our martingale the classical geometric Brownian motion model whereby

$$M_{C}(t) = e^{\sigma W(t) - \frac{\sigma^{2}t}{2}}$$

• Compare this with the discounted approach which would give a discounted shock integral with

$$M_D(t) = \exp\left(\int_0^t e^{-ru}\sigma dW(u) - \frac{\sigma^2}{2}\int_0^t e^{-2ru}du\right)$$

Perpetuity values

• The values of these perpetuities are respectively w_C , w_D with

$$w_{C} = c \int_{0}^{\infty} e^{-rt} N\left(-\frac{\ln(a/S_{0})}{\sigma\sqrt{t}} + \left(\frac{r}{\sigma} - \frac{\sigma}{2}\right)\sqrt{t}\right) dt$$
$$w_{D} = c \int_{0}^{\infty} e^{-rt} N\left(-\frac{\ln(a/S_{0})}{\sigma\sqrt{\frac{1-e^{-2rt}}{2r}}} + \frac{rt}{\sigma\sqrt{\frac{1-e^{-2rt}}{2r}}}\right) dt$$

• For 70% down or $a/S_0 = .7$ at an interest rate of 3% and a volatility of 25% we have

 $w_C = 21.1327c$ $w_D = 29.5941c$

• For 70% down or $a/S_0 = .7$ at an interest rate of 3% and a volatility of 25% we have

 $w_C = 21.1327c$ $w_D = 29.5941c$

• For a 7.5% coupon offered under the classical geometric Brownian motion model the discounted approach would only offer a coupon of 5.43%.

• For 70% down or $a/S_0 = .7$ at an interest rate of 3% and a volatility of 25% we have

 $w_C = 21.1327c$ $w_D = 29.5941c$

- For a 7.5% coupon offered under the classical geometric Brownian motion model the discounted approach would only offer a coupon of 5.43%.
- The classical model could then be seriously underpricing such contracts and thereby offering larger coupons than would be justified by a discounted approach.

Calibrated valuation of time above a barrier

• We calibrate a variance gamma and a discounted variance gamma model to option prices on the S&P500 index as at April 18 2012.

Calibrated valuation of time above a barrier

- We calibrate a variance gamma and a discounted variance gamma model to option prices on the S&P500 index as at April 18 2012.
- For the calibration of the variance gamma model to the 582 options with 30 maturities the results were

σ	=	0.2715
ν	=	6.4921
θ	=	-0.1183
rmse	=	9.2396
aae	=	7.7866
ape	=	0.0596

Calibrated valuation of time above a barrier

- We calibrate a variance gamma and a discounted variance gamma model to option prices on the S&P500 index as at April 18 2012.
- For the calibration of the variance gamma model to the 582 options with 30 maturities the results were

σ	=	0.2715
ν	=	6.4921
θ	=	-0.1183
rmse	=	9.2396
aae	=	7.7866
ape	=	0.0596

• We observe that the *DVG* fit to the surface of option prices is certainly much better.

Dilip B. Madan (Robert H. Smith School of

Price for time above a barrier

• We now wish to evaluate the price of time above a barrier for these two models.

Price for time above a barrier

- We now wish to evaluate the price of time above a barrier for these two models.
- For a barrier of .7 this requires the computation

$$w_{VG} = -\int_{0}^{\infty} \frac{\partial c_{VG}(.7, t)}{\partial K} dt$$
$$w_{DVG} = -\int_{0}^{\infty} \frac{\partial c_{DVG}(.7, t)}{\partial K} dt$$

Price for time above a barrier

- We now wish to evaluate the price of time above a barrier for these two models.
- For a barrier of .7 this requires the computation

$$w_{VG} = -\int_{0}^{\infty} \frac{\partial c_{VG}(.7, t)}{\partial K} dt$$
$$w_{DVG} = -\int_{0}^{\infty} \frac{\partial c_{DVG}(.7, t)}{\partial K} dt$$

• We perform these calculations for the estimated parameters at r = .02966 the rate used in the *DVG* calibrations.

• For these parameters calibrated to SPX options we obtain

 $w_{VG} = 13.2043$ $w_{DVG} = 21.9565$

Dilip B. Madan (Robert H. Smith School of

• For these parameters calibrated to SPX options we obtain

 $w_{VG} = 13.2043$ $w_{DVG} = 21.9565$

• The revised coupon rate would be 4.5104.

• For these parameters calibrated to SPX options we obtain

 $w_{VG} = 13.2043$ $w_{DVG} = 21.9565$

- The revised coupon rate would be 4.5104.
- Coupon rates in markets for similar claims have been between 7 and 10 percent, but comparisons are not that easily made as they also involve credit considerations related to the CDS rate of the issuing entity.

• For these parameters calibrated to SPX options we obtain

 $w_{VG} = 13.2043$ $w_{DVG} = 21.9565$

- The revised coupon rate would be 4.5104.
- Coupon rates in markets for similar claims have been between 7 and 10 percent, but comparisons are not that easily made as they also involve credit considerations related to the CDS rate of the issuing entity.
- We also present a graph of the integrands employed in the calibrated *VG* and *DVG* models.

Numerical estimates

Discounted Value of Probability above 0.7 0.9 0.8 Discounted Probability of being above strike 0.7 DVG 0.6 0.5 VG 0.4 0.3 0.2 0.1 0 20 40 60 80 100 120 140 160 180 200 ٥ Maturity in Years

Figure: The discounted probability of being above the strike of 0.7 for the

Conclusion

calibrated VG and DVG models as a function of the maturity.

Zurich Spring School 2015, April 3

49 / 50

Conclusion I

• It is argued that from the perspective of longer maturities, the discounted stock under no dividend distributions should be modeled as a martingale with a nontrivial limit at infinity.

Conclusion

Conclusion I

- It is argued that from the perspective of longer maturities, the discounted stock under no dividend distributions should be modeled as a martingale with a nontrivial limit at infinity.
- Making an analogy with the discounted gains martingale in the presence of dividends one observes that in the limit the discounted gains martingale converges to the nontrivial limit of discounted gains over all time.
Conclusion

Conclusion I

- It is argued that from the perspective of longer maturities, the discounted stock under no dividend distributions should be modeled as a martingale with a nontrivial limit at infinity.
- Making an analogy with the discounted gains martingale in the presence of dividends one observes that in the limit the discounted gains martingale converges to the nontrivial limit of discounted gains over all time.
- This makes the discounted gains martingale the conditional expectation of its limit.

Conclusion

Conclusion I

- It is argued that from the perspective of longer maturities, the discounted stock under no dividend distributions should be modeled as a martingale with a nontrivial limit at infinity.
- Making an analogy with the discounted gains martingale in the presence of dividends one observes that in the limit the discounted gains martingale converges to the nontrivial limit of discounted gains over all time.
- This makes the discounted gains martingale the conditional expectation of its limit.
- The same should be true for long horizon modeling of balanced stock price processes in the absence of dividend distributions.

Conclusion |

Conclusion I

- It is argued that from the perspective of longer maturities, the discounted stock under no dividend distributions should be modeled as a martingale with a nontrivial limit at infinity.
- Making an analogy with the discounted gains martingale in the presence of dividends one observes that in the limit the discounted gains martingale converges to the nontrivial limit of discounted gains over all time.
- This makes the discounted gains martingale the conditional expectation of its limit.
- The same should be true for long horizon modeling of balanced stock price processes in the absence of dividend distributions.
- Out of balance, and then in the absence of a limiting random variable, *booms* or *busts* are distinguished by the manner in which one loses uniform integrability, by either the positive shock or negative shock martingales failing to be uniformly integrable.

• New models are proposed for the discounted stock price that are stochastic exponentials and conditional expectations of nontrivial random variables at infinity.

- New models are proposed for the discounted stock price that are stochastic exponentials and conditional expectations of nontrivial random variables at infinity.
- A particular example is the discounted variance gamma model for which a closed form for the characteristic function is developed.

- New models are proposed for the discounted stock price that are stochastic exponentials and conditional expectations of nontrivial random variables at infinity.
- A particular example is the discounted variance gamma model for which a closed form for the characteristic function is developed.
- In addition procedures are presented for the valuation of path dependent stochastic perpetuities that never have a payoff and it is argued that as they may be easily valued theoretically, that markets may value as them as well.

- New models are proposed for the discounted stock price that are stochastic exponentials and conditional expectations of nontrivial random variables at infinity.
- A particular example is the discounted variance gamma model for which a closed form for the characteristic function is developed.
- In addition procedures are presented for the valuation of path dependent stochastic perpetuities that never have a payoff and it is argued that as they may be easily valued theoretically, that markets may value as them as well.
- For the new models it is also observed that implied volatility curves do not flatten out as they have nontrivial non-Gaussian limits at infinity.