Achieving near-perfect clustering for high dimension, low sample size data

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AG DANK/BCS Meeting 2013 in London

November 08 - 09, 2013, Department of Statistical Science , University College London, London, the United Kingdom **Room**: the Galton Lecture Theatre, Room 115, 1-19 Torrington Place, 11:15 – 11:35, 09/11/2013 (Sat.)

This work is supported by Grant-in-Aid for JSPS Fellows.

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- Notation and General setting
 - K: Number of clusters, N: Sample size, p: dimensions,
 - n_k : Sample size of Cluster k, $N = \sum_{k=1}^{K} n_k$,
 - $X_k^{(p)}$: *p*-dimensional random vector of Cluster *k* ,

-
$$X_k^{(p)} := (X_{k1}, \ldots, X_{kp})^T, \ d_{ij}^{(p)} := \left\| X_i^{(p)} - X_j^{(p)} \right\|$$

- $X_k^{(p)} \ (k = 1, \ldots, K)$ are independent.

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• Notation and General setting (k = 1, ..., K)

(a)
$$\exists M > 0; \forall s \in \mathbb{N}; \mathbb{E}\left[|X_{ks}|^4\right] < M$$

(b)
$$\frac{1}{p} \sum_{s=1}^{p} \mathbb{E}[X_{ks}]^2 \to \mu_k^2 \quad \text{as} \quad p \to \infty$$

(c)
$$\frac{1}{p} \sum_{s=1}^{p} \operatorname{Var}(X_{ks}) \to \sigma_k^2$$
 as $p \to \infty$

(d)
$$\frac{1}{p} \sum_{s=1}^{p} \left\{ \mathbb{E}[X_{ks}] - \mathbb{E}[X_{ls}] \right\}^2 \to \delta_{kl}^2 \quad \text{as} \quad p \to \infty$$

• Notation and General setting (k = 1, ..., K)

(e)
$$\frac{1}{p} \sum_{s=1}^{p} \mathbb{E}[X_{ks}] \mathbb{E}[X_{ls}] \to \eta_{kl}$$
 as $p \to \infty$

(f) There is some permutation of $X_k^{(\infty)}$,

which is ρ -mixing*.

 \mathbf{n}

*The concepts of ρ -mixing is useful as a mild condition for the development of laws of large number.

• <u>Hall et al. (2005; JRSS B)</u>

– The distance between data vectors from a same cluster is approximately-constant after scaled by \sqrt{p} !



- The distance between data vectors from different clusters is also approximately constant after scaled by \sqrt{p} !



1.2. Difficulty of clustering for HDLSS data

- Hierarchical clustering in HDLSS contexts
 - $\hspace{0.1in} \boldsymbol{U}_1^{(p)}, \hspace{0.1in} \boldsymbol{U}_2^{(p)}, \hspace{0.1in} \boldsymbol{U}_3^{(p)} \overset{\text{i.i.d.}}{\sim} \boldsymbol{X}_1^{(p)}; \hspace{0.1in} \boldsymbol{V}^{(p)} \overset{\text{i.i.d.}}{\sim} \boldsymbol{X}_2^{(p)}$



 $\sqrt{2}\sigma_1 \ge \sqrt{\delta_{12}^2 + \sigma_1^2 + \sigma_2^2}$

- In some cases, classical methods do not work well...

- Maximal data piling (MDP) distance (Ahn and Marron, 2007)
 - The orthogonal distance between the affine subspaces generated by the data vectors in each cluster.



- Clustering with MDP distance (Ahn, et al., 2013)
 - Find successive binary split, each of which creates two clusters in such a way that the MDP distance between them is as large as possible.



- MDP distance Clustering (Ahn, et al., 2013)
 - A sufficient condition for the label consistency

$$\delta_{12}^2 + \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2} > \max\left\{\frac{n_1 + G}{n_1 G}\sigma_1^2, \frac{n_2 + G}{n_2 G}\sigma_2^2\right\}$$

- -where $G \leq \max\{n_1, n_2\}.$
- If $\delta_{12}^2 > 0$ is sufficient large, the label consistency holds.

- Some problems of MDP clustering
 - The sufficient condition depends on variances (and sample sizes.
 - Cannot discover differences between variances in each cluster.

Avoiding stereotypes of clustering method, we can conduct simple and effective methods based on a distance matrix or an inner product matrix.

3.1 Main idea

- Toy example:
 - $X_1 \sim N_p(0, I_p)$
 - For $c \neq 1$, $X_2 \sim N_p(\mathbf{0}, cI_p)$,
 - The condition of Ahn et al. (2013) dose not hold.



3.1 Main idea



Standardized distances converge to some constants in prob. ***Distance** *"vectors"* have the cluster information!*

- Proposed method
 - Step 1. Compute the distance matrix D from the data

matrix X (or the inner product matrix $S := XX^T$).

– Step 2. Calculate the following distances ($\Xi := (\xi_{ij})_{n \times n}$),

$$\xi_{ij} = \sqrt{\sum_{s \neq i, \ s \neq j} (d_{is} - d_{js})^2} \quad \left(\text{ or } := \sqrt{\sum_{t \neq i, \ t \neq j} (s_{it} - s_{jt})^2} \right).$$

-<u>Step 3.</u> For the matrix Ξ , apply a usual clustering method.



Data matrix X







Inner product S

• <u>*K*-means Type</u>

$$Q_{p}(\mathscr{C} \mid K) := \sum_{i=1}^{N} \min_{k} \sum_{j \neq i} \left(d_{ij}^{(p)} - \bar{d}_{kj}^{(p)} \right)^{2},$$

where $\bar{d}_{kj}^{(p)} := \frac{1}{n_{k} - 1} \sum_{i \neq j} d_{ij}^{(p)}.$

– We can optimize this by the usual *k*-means algorithm.

- Important property
 - Under the assumptions a) \sim f), for all $K^* \geq K$,

$$\min_{\mathscr{C}} Q_p(\mathscr{C} \mid K^*) \stackrel{\mathbb{P}}{\longrightarrow} 0 \quad \text{as} \ p \to \infty.$$

• Theoretical results of the *k*-means type - In the case of using a distance matrix

Assume a) \sim f).

If $\forall k, l \ (k \neq l); \ \sigma_k \neq \sigma_l \ \text{or} \ \delta_{kl}^2 > 0$,

then the estimate label vector converges to the true label vector in probability as $p \to \infty$.

- Ahn et al., 2013
$$\delta_{12}^2 + \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2} > \max\left\{\frac{n_1 + G}{n_1 G}\sigma_1^2, \frac{n_2 + G}{n_2 G}\sigma_2^2\right\}$$

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4. Conclusion

- In this presentaion,
 - Introduce geometric representations of HDLSS data,
 - Propose a new efficient clustering method for HDLSS data.
- <u>Remark:</u>
 - In HDLSS contexts,

the closeness between data points may not be meaningful,

but "vectors" of distances have the cluster information!

Reference

[01] Ahn, J., Lee, M.H., and Yoon, Y.J. (2013). Clustering high dimension, low sample size data using the maximal data piling distance, *Statist. Sinica*.

[02] Ahn, J., Marron, J.S., Muller, K.M., and Chi, Y. (2007). The high-dimension, low-sample-size geometric representation holds under mild condition, *Biometrika*, **94** (3), 760–766.

[03] Hall, P., Marron, J.S., and Neeman, A. (2005). Geometric representation of high dimension, low sample size data, J. R. Statist. Soc. B, 67 (3), 427 – 444.

[04] Kolmogorov, A.N. and Rozanov, Y.A. (1960). On strong mixing conditions for stationary Gaussian processes. *Theor. Probab. Appl.* **5**, 204-208.

[05] Sun, W., Wang, J., and Fang, Y. (2012). Regularized k-means clustering of highdimensional data and its asymptotic consistency. *Electron. J. Statist.* **6**, 148–167.

A. Definition of ρ -mixing

• *p*-mixing (Kolmogorov and Rozanov, 1960; Theor. Probab. Appl.)

– For
$$-\infty \leq J \leq L \leq \infty$$
 ,

 \mathcal{F}_J^L : the σ -field of events generated by the r.v.s $(Z_i, J \leq i \leq L)$.

– For any σ -field \mathcal{A} ,

 $L_2(\mathcal{A})$: the space of square-integrable, \mathcal{A} -measurable r.v.s.

– For each $m \ge 1$, define the maximal correlation coefficient

 $\rho(m) := \sup |\operatorname{Corr}(f, g)|, \quad f \in L_2(\mathcal{F}_{-\infty}^j), \ g \in L_2(\mathcal{F}_{j+m}^\infty),$ where $j \in \mathbb{Z}$.

- The sequence $\{Z_i\}$ is said to be ρ -mixing if $\rho(m) \to 0$ as $m \to \infty$.